RESEARCH PAPER

Tobacco smoking and excess mortality in multiple sclerosis: a cohort study

Ali Manouchehrinia,1 Mikail Weston,1 Christopher R Tench,1 John Britton,2 Cris S Constantinescu1

ABSTRACT
Objective As patients with multiple sclerosis (MS) have more than 2.5-fold increased mortality risk, we sought to investigate the impact of tobacco smoking on the risk of premature death and its contribution to the excess mortality in MS patients.

Methods We studied 1032 patients during the period 1994–2013 in a UK-based register. Cox regression model was used to investigate the impact of smoking on the risk of premature death, controlling for confounders. Smoking-specific mortality rates were compared with the UK general population.

Results Of 923 patients with clinically definite MS, 80 (46 males and 34 females) had died by December 2012. HRs for death in current smokers and ex-smokers relative to never smokers were 2.70 (95% CI 1.59 to 4.58, p<0.001) and 1.30 (95% CI 0.72 to 2.32; p = 0.37). The standardised mortality ratio, compared with the UK general population, when stratified by smoking status was 3.83 (95% CI 2.71 to 5.42) in current smokers, 1.96 (95% CI 1.27 to 3.0) in ex-smokers and 1.27 (95% CI 0.87 to 1.86) in non-smokers. Never smokers and ex-smokers with MS had similar mortality rates compared with never smokers and ex-smokers without MS in the male British doctors cohort (1.12 (95% CI 0.63 to 1.97) and 0.54 (95% CI 0.26 to 1.14), respectively), while current smokers with MS had 84% higher rate of death compared with current smokers without MS (95% CI 1.24 to 2.72).

Conclusions Tobacco smoking can account for some of the excess mortality associated with MS and is a risk determinant for all-cause and MS-related death.

INTRODUCTION
Multiple sclerosis (MS) is associated with a range of largely irreversible and progressive disability but is not considered a lethal disease. Life span in patients with MS is estimated to be 5–10 years shorter than in the general population.1–3 Mortality is a major health outcome, which is often ignored in the studies of neurological conditions such as MS,4 possibly due to the long duration of follow-up required. As a result, there is little information about the non-demographic risk factors associated with increased mortality rates. A range of lifestyle, environmental risk factors, demographic characteristics and healthcare interventions influence mortality. Among the lifestyle factors, tobacco smoking is the most prominent preventable cause of premature death,5 being associated with a significant decrease in life expectancy,6–7 and has previously been linked to more severe disease in MS patients.8–11 Tobacco smoking is a complex factor in the aetiology and progression of neurological conditions. While tobacco smoking has been shown to reduce the risk of mortality in patients with PD,12 its effect on the risk of mortality in MS patients is not clear. Higher prevalence of smoking in MS populations13 may explain, in part, the reported shorter life span in individuals with MS compared with the general population.

In this study, we investigated the relationship between tobacco smoking and the life span and mortality rates in a large cohort of MS patients and examined whether smoking can explain the reported increased mortality seen in MS populations.

METHOD
Settings and study design
We studied participants in the Nottingham University Hospitals MS Clinics database during the period 1994–2013. This is a major regional referral centre for MS in East Midlands Counties of England. Our methods and cohort characteristics have been described in detail elsewhere.11 Briefly, the centre covers more than 3000 MS patients. Among all the patients who are regularly seen in these MS specialised clinics, 1032 patients were routinely followed up and details of the disease clinical course, disability scores, date of diagnosis and disease onset, treatment, comorbid conditions, etc., were systematically documented. Our final study population consisted of patients with clinically definite MS according to the McDonald and/or Poser criteria14 15 made by an MS specialist neurologist.

Measurements
Clinical data and smoking history
Smoking history was obtained during patients’ first clinic visit at the time of disease onset and/or diagnosis, and patients were grouped as non-smokers, ex-smokers or current smokers. In the majority of the cases, smoking history was updated and recorded more than once after the disease onset during regular clinic follow-ups. For this study, we used the latest smoking status recorded in the database. The date of the first manifestation of the disease, date of diagnosis, duration of exposure to disease-modifying treatments (DMTs), latest Expanded Disability Status Scale (EDSS) score16 recorded in clinic, sex and initial clinical course of the disease (relapsing-remitting (RR) vs primary-progressive (PP)) were used in this report. We also calculated the global MS severity score (MSSS), which integrates EDSS...
score and disease duration according to the guidelines published by Roxburgh et al.\textsuperscript{17}

Vital status and cause of death

Vital status of patients was monitored and the exact date of death was obtained through linkage of the MS cohort to National Health Service (NHS) vital statistics as of December 2012 (index date). For the cause of death, both medical records and death certificates were used. Causes of death were categorised into MS-related (death due to MS disability such as bronchopneumonia, pulmonary embolism, sepsis) and non-MS-related deaths (eg, cancers, suicides, cardiovascular diseases).\textsuperscript{18}

Statistical analysis

Patient data from the time of entry into the database until either death or the index date (December 2012), whichever occurred first, were included. Time at risk and death rates were calculated per 1000 person-years. CI for median survival age could not be estimated as less than half the patients have died by December 2012. Since the distribution was reasonably normal, we used parametric measures and calculated mean survival from birth and after MS onset. Smoking-specific rates were calculated using Cox proportional hazard regression models controlled for sex, onset age, use of DMTs (in a binary group of ≥1 year or <1 year) and initial clinical course of MS (RR vs PP). By basing the survival time on age, we made sure that the effects of age as one of the most important factors in progression have been properly controlled for. The final model was checked for proportionality assumption based on the Schoenfeld residuals, and stratification was made if necessary to hold the proportionality assumption. We also calculated standardised mortality ratios (SMRs) and performed external comparison of our mortality to both the UK general population data and mortality rates of male British doctors cohort.\textsuperscript{19} The male British doctors study contains the mortality rates stratified by age and smoking status. The smoking-specific mortality rates from this cohort enabled us to measure the excess mortality in different smoking categories in male patients with MS. SMR was calculated by dividing the observed number of deaths in the cohort by the number of deaths expected from the general population for each sex and age band stratum (20–24, 25–29, ..., up to >85 years). Since all the deaths (n=80) in our cohort occurred in the 2001–2012 calendar period, we used 2006 UK mortality rates as the corresponding reference rates, obtained from the UK Office for National Statistics for each sex and age group.\textsuperscript{20} All statistical analyses were performed with Stata V11.2.\textsuperscript{21}

FINDINGS

Our cohort consisted of 923 (89% of 1032 subjects) individuals with clinically definite MS and full data for analysis. In total, 70% of the subjects were female, 11% had PP MS at disease onset and 89% had RR MS at onset. Of those RR MS subjects, 40% had transitioned to secondary progressive (SP) MS. In general, deceased patients in our cohort had significantly older onset age, longer disease duration and higher MSSS and EDSS scores compared with survivors. Also, 58% of the deceased patients and 30% of survivors were male; 45% of the survivors had received DMTs for 1 year or longer compared with only 16% of deceased subjects. Average treatment duration was 4.73 (SD ±3.36) years for the whole cohort. Despite the decrease in the prevalence of ever-smoking in the UK general population,\textsuperscript{22} there was an increase of approximately 12% in the proportion of ever smokers in our cohort in 2000–2010 compared with that of the proportion before 1990 (data not shown) in the MS cohort. Table 1 shows the baseline demographic and MS-specific characteristics of our cohort by vital status at our index date.

Length of follow-up and smoking status

The 923 patients contributed a total of 18 717 person-years of data. There were 80 deaths recorded in our cohort, representing a mortality rate of 4.35 (95% CI 3.43 to 5.34) per 1000 person-years. This rate was higher in males (8.68, 95% CI 6.52 to 11.56) than in females (2.47, 95% CI 1.76 to 3.48). When compared with the frequency of 50% in the whole cohort, 67% of the deceased subjects were ever smokers (n=54). All nine patients with RR MS who died were ever smokers compared with 44% of PP MS and almost 68% of SP MS patients. Of patients with relapse onset MS, 37% among those who were still alive and 67% in the deceased group had transitioned to SP MS.

Survival rates and role of smoking

Estimated mean survival was 76.8 years (95% CI 74.6 to 79) for the whole cohort. Estimated mean survival was 81 (95% CI 78 to 83.6) years in non-smokers, 78.4 (95% CI 75.2 to 81.5) years in ex-smokers and 71.5 (95% CI 68.8 to 74.2) in current smokers (Log-rank test, χ² (2)=20.12, p=0.001) (figure 1A,B). The estimated mean survival was not different between RR MS and PP MS (Log-rank test, χ² (1)=0.38, p=0.53).

While controlling for onset age, DMT and disease initial course and stratifying for sex (to maintain the proportionality assumption), current smokers but not ex-smokers were at higher risk of death, with a HR relative to never smokers of 2.70 (95% CI 1.59 to 4.58; p<0.001) and 1.30 (95% CI 0.72 to 2.32; p=0.37), respectively. Exposure to DMT for 1 year or longer was associated with 50% (HR 0.50, 95% CI 0.26 to 0.93; p=0.03) reduction in the risk of death, and each year increase in age at the onset of MS was associated with 4% (95% CI 0.94 to 0.98; p=0.004) decrease in the risk of death. Risk of death was not significantly increased in PP MS patients compared with RR MS (HR 1.16, 95% CI 0.65 to 1.06; p=0.65).

Since both EDSS and year of birth are significant contributory factors in the risk of death, they were added to the model to check whether the effect size of smoking remains significant.

<table>
<thead>
<tr>
<th>Table 1</th>
<th>General demographic characteristics of the cohort by vital status at December 2012</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Deceased (n=80)</td>
</tr>
<tr>
<td>Age (mean (SD))</td>
<td>56 (±12)</td>
</tr>
<tr>
<td>Sex (female %)</td>
<td>34 (42%)</td>
</tr>
<tr>
<td>Last recorded EDSS (median (range))</td>
<td>7.5 (6.5–8.0)</td>
</tr>
<tr>
<td>MSSS (mean (SD))</td>
<td>7.9 (±2.23)</td>
</tr>
<tr>
<td>Disease duration (median (range))</td>
<td>18 (10–26)</td>
</tr>
<tr>
<td>Type of MS</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Relapsing-remitting</td>
<td>9 (11%)</td>
</tr>
<tr>
<td>Primary-progressive</td>
<td>16 (20%)</td>
</tr>
<tr>
<td>Secondary-progressive</td>
<td>55 (69%)</td>
</tr>
<tr>
<td>DMT ≥1 year</td>
<td>13 (16%)</td>
</tr>
</tbody>
</table>

DMT, disease-modifying treatment; EDSS, expanded disability status scale; MSSS, multiple sclerosis severity score.
When the latest EDSS score and year of birth were added to the model, the effect size of smoking increased to 2.92 (95% CI 1.50 to 5.71; p=0.002) in current smokers and 1.62 (95% CI 0.79 to 3.34; p=0.18) in ex-smokers compared with non-smokers and exposure to DMT for at least a year was associated with 68% reduction in the risk of death (HR 0.32, 95% CI 0.14 to 0.72; p=0.006).

Limiting our survival analysis to those who died from MS-related causes, current smokers were still at higher risk for death, with a HR of 2.93 (95% CI 1.48 to 5.76; p<0.001) relative to non-smokers. Among ex-smokers, the HR was 1.18 (95% CI 0.53 to 2.61; p=0.67).

Disability status and cause of death
Latest recorded disability status was significantly higher in deceased subjects relative to survivors as measured by EDSS score and MSSS. We could not obtain causes of death in two of the decedents. Cause of death was MS-related in 60% of cases (table 2), with bronchopneumonia accounting for the majority of MS-related deaths. Cancer (17.5%) was the commonest non-MS-related cause, and lung cancer the dominant cancer (six deaths). Deaths from cardiovascular diseases (11 cases), suicide (1 case), motor neuron disease (1 case), kidney failure (2 cases), liver cirrhosis and gastrointestinal infarction were also recorded.

**Table 2** Summaries of causes of death, onset age, disability score and disease severity by patients’ lifelong smoking status

<table>
<thead>
<tr>
<th>Cause of death</th>
<th>Ever smoked</th>
<th>Never smoked</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>n</td>
<td>Onset age (median)</td>
</tr>
<tr>
<td>MS-related cause</td>
<td>32</td>
<td>36</td>
</tr>
<tr>
<td>Pneumonia</td>
<td>21</td>
<td>8</td>
</tr>
<tr>
<td>Sepsis</td>
<td>4</td>
<td>8</td>
</tr>
<tr>
<td>Other or not specified</td>
<td>7</td>
<td>8</td>
</tr>
<tr>
<td>Cancer</td>
<td>10</td>
<td>33</td>
</tr>
<tr>
<td>Cardiovascular</td>
<td>7</td>
<td>38</td>
</tr>
<tr>
<td>Other</td>
<td>2</td>
<td>27.5</td>
</tr>
</tbody>
</table>

EDSS, expanded disability status scale; MS, multiple sclerosis.

**Table 3** Standardised mortality ratio for all-cause mortality compared with the UK general population and male British doctors cohort

<table>
<thead>
<tr>
<th></th>
<th>SMR (general population)</th>
<th>SMR (male British doctors)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Overall</td>
<td>1.99 (1.70 to 2.33)</td>
<td>–</td>
</tr>
<tr>
<td>Male</td>
<td>2.41 (1.95 to 2.96)</td>
<td>–</td>
</tr>
<tr>
<td>Female</td>
<td>1.80 (1.40 to 2.30)</td>
<td>–</td>
</tr>
<tr>
<td>Non-smokers</td>
<td>1.27 (0.87 to 1.86)</td>
<td>1.12 (0.63 to 1.97)</td>
</tr>
<tr>
<td>Ex-smokers</td>
<td>1.96 (1.27 to 3.0)</td>
<td>0.54 (0.26 to 1.14)</td>
</tr>
<tr>
<td>Current smokers</td>
<td>3.83 (2.71 to 5.42)</td>
<td>1.84 (1.24 to 2.72)</td>
</tr>
</tbody>
</table>

SMR, standardised mortality ratio and CI5 (95%).

**Figure 1** Cox proportional hazard regression graph showing survival from birth (A) by lifelong smoking status and (B) latest smoking status.

**DISCUSSION**
In this MS population, current smoking was associated with more than 2.5-fold increased risk of death. Current smokers with MS had a reduction of about 10 years in their life expectancy relative to non-smokers with MS. When compared with the general population, increased mortality in this cohort as measured by SMR was seen in current and ex-smokers, but our results indicated that mortality may not be increased in non-smokers with MS. The risk of death in ex-smokers was considerably lower than that in current smokers. When the mortality in...
each smoking category in our patients was compared with the mortality in smoking categories in the British doctors cohort study, we found that current smokers with MS have increased mortality risk (84%) compared with their counterparts without MS. Interestingly, non-smokers and ex-smokers with and without MS had similar mortality rates compared with non-smokers and ex-smokers without MS. These data are consistent with our recent findings of the negative impact of current smoking on disability and severity of the disease in MS.

The estimated survival age of 77 for our MS population in the present study is comparable to the findings from a recent study in the UK and from Canada. Our patients were two times more likely to die prematurely compared with their age-matched and sex-matched counterparts in the UK general population. Although SMRs cannot be comfortably compared due to differences in distribution of the standardisation variables (age and sex) across the study populations, our presented SMR here is lower, though still comparable, with those reported previously. The overall reported SMR here compared with the UK general population is higher than the reported SMR of 1.3 in a French study. Apart from methodological variations, treatment and access to healthcare resources may further explain the lower SMR in the cohort presented here. While 57% of the patients in the French study and 42% of our cohort have been exposed to treatment, the majority of the patients in British Columbia were never exposed to treatment and the follow-up time in the Danish study predated treatment. This may explain some of the differences seen between our reported SMR here and those reported previously.

In our cohort, the risk of death was halved for patients who had been exposed to at least 1 year of DMTs. However, the assignment of DMTs to a selective population of patients may introduce bias to the results. Our results presented here are in line with the findings from the long-term follow-up of patients participating in the original interferon-β clinical trial, which has shown that exposure to treatment may decrease risk of premature death. Also, follow-up of patients in Taiwan showed a reduced risk of mortality in patients with history of exposure to DMTs. The mechanism by which DMTs can reduce the risk of premature death is unclear. We could not observe any influence of DMTs in reducing the risk of reaching two EDSS score milestones 4 and 6 in our cohort (data not shown). One possible explanation is the ability of interferon to reduce the risk of infection. Nevertheless, the potential influence of treatment intervention in MS on mortality deserves further investigation.

Mortality is a major component of population change and many health conditions and interventions reflect on mortality rates. This, along with unique characteristics such as exact date and time, makes mortality a relevant outcome in studies of MS. In the UK, it has been shown that current smokers with MS have 6.7-fold (95% CI 4.16 to 10.87) increased mortality rate compared with the sex-matched and age-matched counterparts without MS while the rate for never smokers was 2.85 (95% CI 1.86 to 4.37). Our results confirm a higher mortality rate in current smokers versus non-smokers (HR 2.7, 95% CI 1.59 to 4.58; p<0.001).

Our study had some limitations. First, it is possible that our study has missed some of the patients with very severe disease who have died before the data collection. Second, our calculated SMR, which used reference mortality rates from the male British doctors study, were limited to the male patients while 70% of our cohort were female. In addition, the reported mortality rates in that study of male British doctors, which is a historic cohort (1951–2001), are higher from the more recent ones. Hence, it is likely that the SMR versus male doctors substantially underestimates the excess mortality for the MS cohort for the non-smoker and current smoker groups. We cannot rule out the possibility of reverse causality bias (or the healthy smoker bias) when ex-smokers probably quit smoking because of the worsening of their MS. As a consequence, the effect size of current smoking versus ex-smokers may be biased by selection of those current smokers who have experienced fewer disabilities, are still alive and continued to smoke.

To our knowledge, this is the first study investigating smoking as a factor influencing life expectancy in patients with MS. Here we have presented data from a cohort of MS patients in the UK, but the relevance of this study is likely to be worldwide. As shown here, smoking is a significant risk factor for all-cause and MS-related death, causing almost 10% years of life expectancy in patients with MS. In contrast, our data suggest that mortality may not be increased substantially in non-smokers with MS relative to the general population and non-smokers in the male British doctors study.

In the general population, tobacco smoking reduces life expectancy mainly by increasing the risk of cancer and cardiovascular diseases. In MS, vascular comorbidities were shown to associate with a more severe disease evolution, which can potentially mediate the effect of smoking on mortality in MS. However, our risk estimates remained almost unchanged when only those who died from MS-related causes were analysed. In addition, we did not observe any increased mortality rate due to cancer and cardiovascular diseases compared with the UK general population (SMR due to cancer 0.61, 95% CI 0.37 to 1.01; and SMR due to cardiovascular diseases 0.83, 95% CI 0.46 to 1.50). The exact mechanism in which smoking increases the risk of premature death in MS deserves further investigation. One possible explanation is the increased risk of mortality due to pneumonia, for which smoking has been shown to be a significant risk factor. In our cohort, the number of ever smokers who died from pneumonia was more than two times higher than never smokers (n=21 vs n=9).

On the basis of our findings, there is a strong association between tobacco smoking and premature mortality in patients with MS. The evidence of reduced risk of death among former smokers compared with current smokers calls for effective cessation strategies. Our data suggest that reduction in prevalence of tobacco smoking in people with MS is one factor that can potentially decrease the excess mortality rates seen in MS patients. Tools and information to help smokers with MS stop smoking and non-smokers to never start smoking should be provided to all patients with MS.

Contributors AM was involved in study concept and design, analysis and interpretation of the data and critical revision of the manuscript for important intellectual content. MW was involved in acquisition of data. CRT was involved in developing the study concept and design, and critical revision of the manuscript for important intellectual content. CSC also supervised the study and was also involved in acquisition of data. MF was involved in analysis and interpretation of results, and critical revision of the manuscript for important intellectual content. JF performed the critical revision of the manuscript for important intellectual content. SC was involved in developing the study concept and design, and critical revision of the manuscript for important intellectual content. CSC also supervised the study and was also involved in acquisition of data.

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Competing interests CSC has received research support, support for travel to scientific meetings, consultancy and speaker honoraria from Biogen Idec, Bayer-Schering, Centocor, GW Pharmaceuticals, Merck-Serono, Morphosys, and Teva Pharmaceuticals.

Ethics approval The study was approved by the National Research Ethics Service East Midlands Ethics Committee Derby-1.

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